PROPHYLACTIC IMPLANTATION OF A DEFIBRILLATOR IN PATIENTS WITH MYOCARDIAL INFARCTION AND REDUCED EJECTION FRACTION

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ABSTRACT

Background Patients with reduced left ventricular function after myocardial infarction are at risk for life-threatening ventricular arrhythmias. This randomized trial was designed to evaluate the effect of an implantable defibrillator on survival in such patients.

Methods Over the course of four years, we enrolled 1232 patients with a prior myocardial infarction and a left ventricular ejection fraction of 0.30 or less. Patients were randomly assigned in a 3:2 ratio to receive an implantable defibrillator (742 patients) or conventional medical therapy (490 patients). Invasive electrophysiological testing for risk stratification was not required. Death from any cause was the end point.

Results The clinical characteristics at base line and the prevalence of medication use at the time of the last follow-up visit were similar in the two treatment groups. During an average follow-up of 20 months, the mortality rates were 19.8 percent in the conventional-therapy group and 14.2 percent in the defibrillator group. The hazard ratio for the risk of death from any cause in the defibrillator group as compared with the conventional-therapy group was 0.69 (95 percent confidence interval, 0.51 to 0.93; P=0.016). The effect of defibrillator therapy on survival was similar in subgroup analyses stratified according to age, sex, ejection fraction, New York Heart Association class, and the QRS interval.

Conclusions In patients with a prior myocardial infarction and advanced left ventricular dysfunction, prophylactic implantation of a defibrillator improves survival and should be considered as a recommended therapy. (N Engl J Med 2002;346:877-83.)

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*The investigators who participated in the Multicenter Automatic Defibrillator Implantation Trial II (MADIT-II) are listed in the Appendix.
METHODS

Organization of the Trial

The trial began on July 11, 1997, and enrolled patients from 76 hospital centers (71 in the United States and 5 in Europe). The protocol was approved by the institutional review boards of the participating hospitals. A data and safety monitoring board independently reviewed the results at regular intervals throughout the trial. All investigators agreed to abide by the conflict-of-interest guidelines described by Healy et al.4 A description of the design and study protocol has been published previously.5

Recruitment and Follow-up

Patients of either sex who were more than 21 years of age (there was no upper age limit) were eligible for the study if they had had a myocardial infarction one month or more before entry, as documented by the finding of an abnormal Q wave on electrocardiography, elevated cardiac-enzyme levels on laboratory testing during hospitalization for suspected myocardial infarction, a fixed defect on thallium scanning, or localized akinesis on ventriculography with evidence of obstructive coronary disease on angiography, and an ejection fraction of 0.30 or less within three months before entry, as assessed by angiography, radionuclide scanning, or echocardiography. Potentially eligible patients were referred by local cardiologists, internists, and primary care physicians. Patients were not required to undergo electrophysiological screening for inducible ventricular arrhythmias.

Patients were excluded from enrollment if they had an indication approved by the Food and Drug Administration (FDA) for an implantable defibrillator; were in New York Heart Association functional class IV at enrollment; had undergone coronary revascularization within the preceding three months; had had a myocardial infarction within the past month, as evidenced by measurement of cardiac-enzyme levels; had advanced cerebrovascular disease; were of childbearing age and were not using medically prescribed contraceptive measures; had any condition other than cardiac disease that was associated with a high likelihood of death during the trial; or were unwilling to sign the consent form for participation.

When the trial began in July 1997, eligible patients had to have frequent or repetitive ventricular ectopic beats during 24-hour Holter monitoring. On January 1, 1998, after the enrollment of 23 patients, the executive committee eliminated this requirement because almost all eligible patients had such arrhythmias. On May 4, 2001, the executive committee increased the enrollment goal from 1200 to 1500 patients, so that enrollment would be ongoing while data on outcomes were still accruing.

Randomization

After patients had provided written informed consent, a base-line clinical history and 12-lead electrocardiogram were obtained and a physical examination was conducted. The patients were randomly

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Figure 1. Sequential Monitoring in the Triangular Design.

The log-rank statistic is a measure of the cumulative differences in survival between the two treatments. A positive value for the log-rank statistic indicates that treatment with the defibrillator was superior to conventional medical therapy. The variance of the log-rank statistic is closely related to the number of deaths. Three boundaries are shown: one indicating the inefficacy of the implantation of a defibrillator as compared with conventional medical therapy, one indicating no difference between the groups, and one indicating the superiority of treatment with a defibrillator. The trial was stopped on November 20, 2001, shortly after the upper boundary, indicating superiority of the defibrillator, was reached (P = 0.027). The values of the log-rank statistic continue to increase after the termination of the trial (P = 0.016) as a result of a lag in reporting additional survival data and five additional deaths that occurred before the stopping date but were uncovered during the closeout procedure (on January 16, 2002).
assigned in a 3:2 ratio to receive either an implantable defibrillator or conventional medical therapy. Patients who were assigned to the defibrillator group were not responsible for the costs of the defibrillator, implantation, or the hospitalization for the procedure.

Therapy

Transvenous defibrillator systems (Guidant, St. Paul, Minn.) that had been approved by the FDA were used in the trial. Standard techniques were used to implant the defibrillators. The defibrillators were tested during the implantation procedure, and every effort was made to achieve defibrillation within a 10-J safety margin. Programming the defibrillator and prescribing medications were left to the discretion of the patients' physicians. The appropriate use of beta-blockers, angiotensin-converting–enzyme inhibitors, and lipid-lowering drugs was strongly encouraged in both study groups.

Statistical Analysis

The primary end point was death from any cause. Analysis was performed according to the intention-to-treat principle. The trial was designed to have 95 percent power to detect a 38 percent reduction in the two-year mortality rate among the patients in the defibrillator group, given a postulated two-year mortality rate of 19 percent among patients assigned to conventional therapy, with a two-sided significance level of 0.05. For proportional-hazards modeling, power was maintained for a true hazard ratio of 0.63, after allowance for crossovers. We used a triangular sequential design, which was modified for two-sided alternatives and corrected for the lag in obtaining data accrued but not reported before the termination of the trial, for weekly monitoring, with preset boundaries to permit termination of the trial if the defibrillator therapy was found to be superior to, inferior to, or equal to conventional medical therapy. Secondary analyses were performed with use of the Cox proportional-hazards regression model. We determined survival curves according to the method of Kaplan and Meier, with comparisons of cumulative mortality based on logarithmic transformations. P values were termed nominal when they were not adjusted for sequential monitoring. All P values were two-tailed.

At the recommendation of the data and safety monitoring board the trial was stopped on November 20, 2001, shortly after an analysis revealed that the difference in mortality between the two groups had reached the prespecified efficacy boundary (P = 0.027) (Fig. 1). Analyses used version 2.0 of the data base, which was released on January 16, 2002. The investigators had full access to the data and performed the data analysis with no limitations imposed by the sponsor.

RESULTS

Study Population

The clinical characteristics of the 1232 randomized patients are provided in Table 1. Follow-up averaged 20 months (range, 6 days to 53 months). The base-line characteristics and the prevalence of the use of various cardiac medications at the time of the last follow-up visit were similar in the two groups. Fifty-four crossovers occurred. Twenty-two patients in the conventional-therapy group (4.5 percent) received a defibrillator during the trial, 21 for documented or suspected malignant ventricular arrhythmias and 1 at the physician’s discretion. Twenty-one patients assigned to the defibrillator group (2.8 percent) did not have a defibrillator implanted, and 11 had their defibrillator removed during the trial (1.5 percent), including 9 who underwent heart transplantation. Twelve patients had their defibrillator deactivated during the trial, usually as a result of terminal illness. There were 8749 scheduled follow-up visits during the trial, with a 94 percent rate of attendance in the conventional-therapy group and a 97 percent rate in the defibrillator group. The status of three patients was not known at the termination of the trial (one patient in the conventional-therapy group and two in the defibrillator group). All three patients were known to be alive within six months before the trial ended.

**Table 1. Clinical Characteristics of the 1232 Patients.**

<table>
<thead>
<tr>
<th>Characteristic</th>
<th>Defibrillator Group (N=742)</th>
<th>Conventional-Therapy Group (N=490)</th>
</tr>
</thead>
<tbody>
<tr>
<td>Age (yr)</td>
<td>64±10</td>
<td>65±10</td>
</tr>
<tr>
<td>Male sex (%)</td>
<td>84</td>
<td>85</td>
</tr>
<tr>
<td>NYHA functional class (%)†</td>
<td>35 35 25 5 4 53 53 33 80 58 45 88 87</td>
<td></td>
</tr>
<tr>
<td>Blood uric acid &gt;25 mg/dl (8.92 mmol/liter)</td>
<td>29 32</td>
<td>29 32</td>
</tr>
<tr>
<td>Arrtial fibrillation (%)</td>
<td>9 8</td>
<td>9 8</td>
</tr>
<tr>
<td>QRS interval &gt;0.12 sec (%)</td>
<td>50 51</td>
<td>50 51</td>
</tr>
<tr>
<td>Nonspecific conduction defect (%)</td>
<td>22 26</td>
<td>22 26</td>
</tr>
<tr>
<td>Right bundle-branch block (%)</td>
<td>9 7</td>
<td>9 7</td>
</tr>
<tr>
<td>Left bundle-branch block (%)</td>
<td>19 18</td>
<td>19 18</td>
</tr>
<tr>
<td>Left ventricular ejection fraction (%)</td>
<td>23.5 23.5</td>
<td>23.5 23.5</td>
</tr>
<tr>
<td>Medications at last contact (%)‡</td>
<td>Amiodarone 13 10</td>
<td>Amiodarone 13 10</td>
</tr>
<tr>
<td>Angiotensin-converting–enzyme inhibitors (%)</td>
<td>68 72</td>
<td>68 72</td>
</tr>
<tr>
<td>Beta-blockers (%)</td>
<td>70 70</td>
<td>70 70</td>
</tr>
<tr>
<td>Calcium-channel blockers (%)</td>
<td>9 9</td>
<td>9 9</td>
</tr>
<tr>
<td>Class I antiarrhythmic agents (%)</td>
<td>3 2</td>
<td>3 2</td>
</tr>
<tr>
<td>Digitalis (%)</td>
<td>57 57</td>
<td>57 57</td>
</tr>
<tr>
<td>Diuretics (%)</td>
<td>72 81</td>
<td>72 81</td>
</tr>
<tr>
<td>Lipid-lowering statin drugs (%)</td>
<td>67 64</td>
<td>67 64</td>
</tr>
</tbody>
</table>

*Plus–minus values are means ±SD. Patients were randomly assigned in a 3:2 ratio to receive an implanted defibrillator (60.2 percent) or conventional medical therapy (39.8 percent).

†Values reflect the highest New York Heart Association (NYHA) functional class recorded in the three-month period before enrollment. Eligibility was limited to patients who were in NYHA class I, II, or III at the time of enrollment.

‡The mean interval from enrollment to the time of the last follow-up visit in which medication use was recorded was 17 months in the conventional-therapy group and 18 months in the defibrillator group.
Outcome

The sample path of the sequential trial is presented in Figure 1. The trajectory of the path indicates the superiority of defibrillator therapy over nondefibrillator therapy, with \( P = 0.016 \) (two-sided, adjusted for the stopping rule).

The numbers of deaths in each group and the hazard ratio for death per unit of time are presented in Table 2. The hazard ratio of 0.69 indicates a 31 percent reduction in the risk of death at any interval among patients in the defibrillator group as compared with patients in the conventional-therapy group.

### Table 2. Hazard Ratio for Death.

<table>
<thead>
<tr>
<th>VARIABLE</th>
<th>DEFIBRILLATOR GROUP (N=742)</th>
<th>CONVENTIONAL-THERAPY GROUP (N=490)</th>
<th>HAZARD RATIO (95% CI) *</th>
<th>P VALUE †</th>
</tr>
</thead>
<tbody>
<tr>
<td>No. of deaths (%)</td>
<td>105 (14.2)</td>
<td>97 (19.8)</td>
<td>0.69 (0.51–0.93)</td>
<td>0.016</td>
</tr>
</tbody>
</table>

*The value is the risk of death per unit of time among patients who were randomly assigned to the defibrillator group as compared with that among patients who were randomly assigned to the conventional-therapy group. The hazard ratio, 95 percent confidence interval (CI), and \( P \) value were derived from the sequential design with use of a proportional-hazards assumption and took into account the sequential stopping rule.

†The \( P \) value is two-sided.

Kaplan-Meier estimates of survival in the two groups are shown in Figure 2. The two survival curves began to diverge at approximately nine months and continued their separate paths thereafter (nominal \( P = 0.007 \)). These survival curves represent reductions in the rates of death after defibrillator therapy of 12 percent at one year (nominal 95 percent confidence interval, \( -27 \) to 40 percent), 28 percent at two years (nominal 95 percent confidence interval, \( 4 \) to 46 percent), and 28 percent at three years (nominal 95 percent confidence interval, \( 5 \) to 46 percent). There were no significant differences in the effect of defibrillator therapy on survival in subgroup analyses stratified according to age, sex, ejection fraction, New York Heart Association class, or the QRS interval (Fig. 3). There were also no significant differences in the effect of defibrillator therapy on survival in subgroup analyses classified according to the presence or absence of hypertension, diabetes, left bundle-branch block, or atrial fibrillation; the interval since the most recent myocardial infarction (six months or less vs. more than six months); the type of defibrillator implanted (single chamber vs. dual chamber); or the blood urea nitrogen level (25 mg per deciliter or less vs. more than 25 mg per deciliter).

Adverse Effects

Serious complications related to defibrillator therapy were infrequent. No deaths occurred during im-
plantation. Thirteen lead problems (1.8 percent) and five nonfatal infections (0.7 percent) required surgical intervention in the defibrillator group. The incidence of new or worsened heart failure was slightly higher in the defibrillator group than in the conventional-therapy group. Specifically, 73 patients (14.9 percent) in the conventional-therapy group and 148 in the defibrillator group (19.9 percent) were hospitalized with heart failure, representing 9.4 and 11.3 patients so hospitalized per 1000 months of active follow-up, respectively (nominal P=0.09).

**DISCUSSION**

Our findings indicate that the implantation of a defibrillator improves survival among patients with a prior myocardial infarction and a left ventricular ejection fraction of 0.30 or less. As compared with conventional medical therapy, defibrillator therapy was associated with a 31 percent reduction in the risk of death. Electrophysiologic testing or inducible ventricular arrhythmias were not eligibility criteria. The two groups were balanced and received appropriate therapy with standard cardiac agents; high percentages of patients in both groups received angiotensin-converting–enzyme inhibitors, beta-blockers, diuretics, and lipid-lowering statin drugs.

In contrast with the earlier Multicenter Automatic Defibrillator Implantation Trial, in which the survival rate improved within the first few months after the implantation of the device, in the current study the survival benefit began approximately nine months after the device was implanted. The difference may reflect a somewhat lower mortality rate in the conventional-therapy group in the current study, the absence of risk stratification for arrhythmia as an entry criterion, the use of a lower cutoff value for the ejection fraction as a criterion for eligibility, and the use of more vigorous medical treatment. These same pop-
ulation differences may also explain the relatively higher hazard ratio in our study than in the earlier study (0.69 vs. 0.46). We are concerned by the finding that new or worsened heart failure requiring hospitalization was slightly more frequent in the defibrillator group than in the conventional-therapy group. Patients saved from malignant ventricular arrhythmias by the implantation of a defibrillator live longer than conventionally treated patients and would thus have more time for heart failure to develop. Defibrillator shocks might contribute to rehospitalization and myocardial injury. Backup ventricular pacing may impair ventricular function. Patients with a reduced ejection fraction who receive an implantable defibrillator should be carefully monitored for the development or exacerbation of heart failure.

An estimated 3 million to 4 million patients have coronary heart disease and advanced left ventricular dysfunction in the United States, and there are approximately 400,000 new cases annually. If a meaningful number of these patients receive an implantable defibrillator prophylactically, the cost to the health care system would be substantial. We hope that market forces will drive down the cost of this therapy.

Our findings show that the implantation of a defibrillator improves survival in patients with a prior myocardial infarction and advanced left ventricular dysfunction. Thus, prophylactic implantation of a defibrillator is recommended in these patients.

Supported by a research grant from Guidant, St. Paul, Minn., to the University of Rochester School of Medicine and Dentistry. Drs. Cannom, Daubert, and Higgins have given lectures sponsored by Guidant.

We are indebted to the patients who participated in this trial, to the study coordinators at each of the enrolling centers, and to Guidant Corporation for its support, sustained commitment, and the independence it provided to the investigators in the conduct of this study.

APPENDIX

The following investigators also participated in the MADIT-II (listed in descending order of the number of randomly assigned patients):

Investigators — L. Berenbom, Mid-America Research Organization and University of Kansas Medical Center Research Institute, Kansas City; D. Glascock, Toledo Hospital, Toledo, Ohio; H. Klein, University Hospital Magdeburg, Magdeburg, Germany; H. Pirschnier, Klinikum Clinic, Bad Nauheim, Germany; J. Butler, Florida Heart Association, Fort Myers; J. Daubert, Brigham and Women's Hospital, Boston; G. Sosa-Suarez, St. Peter's Hospital, Albany, N.Y.; D. Wilkinson, Swedish Medical Center–Providence Campus, Seattle; D. Fitzgerald, Wake Forest University, Winston-Salem, N.C.; Y. Greenberg, Mainzmedizin Medical Center, Bysslín, N.Y.; T. Talbert, Park Ridge Medical Center, Chattanooga, Tenn.; R. Jadonath, North Shore University Hospital, Manhasset, N.Y.; J. Windle, University of Nebraska Medical Center, Omaha; L. Siddoway, York Hospital, York, Pa.; D. Borsovski, Consultant in Cardiology, Inc., Erie, Pa.; E. Nash, Peninsula Cardiology Associates, Salisbury, Md.; S. Geraci, University of Florida, Jacksonville; S. Brownstein, St. Vincent Mercy Medical Center, Toledo, Ohio; M. Schali, Leiden University Hospital, Leiden, the Netherlands; A. Medina, Bikur Cholim Hospital, Jerusalem, Israel; E. Platia, Washington Hospital Center, Washington, D.C.; P. Friedman, Cardiovascular Specialists, Hyannis, Mass.; J. Zebede, Mount Sinai Medical Center, Miami; E. Leonelli, Cardiology Associates of Mobile, Mobile, Ala.; E.W. Grogan, Wisconsin Heart, Madison; W. Crawford, Montgomery Cardiovascular Associates, Montgomery, Ala.; D. Wattoo, Heart Center of Nevada, Las Vegas; J. Souza, Asheville Cardiology Associates, Asheville, N.C.; C. Fuenzalida, Western Cardiology Associates, Aurora, Colo.; S. Lanza, Florida Heart Group, Orlando; B. Bel hassan, Tél-Aviv Medical Center, Tel Aviv, Israel; R. Winkle, Sequoia Hospital, Palo Alto, Calif.; M. Illovsky, National Naval Medical Center, Bethesda, Md.; A. Dougherty, University of Texas, Houston; O. Fujimura, State University of New York Upstate Medical University; S. Scardone, University of Iowa Health Care, Iowa City; C. Rizo-Patron, Cardiology Associates of Lubbock, Lubbock, Tex.; D. Pederson, Austin Heart, Austin, Tex.; F. Gilliam, Virginia Cardiovascular Specialists, Richmond; M. Brown, University of Florida, Jacksonville; S. Compton, Alaska Research Foundation, Greensboro, N.C.; J. Merillat, Valley Heart Associates, Oklahoma City; A. Leonelli, Cardiology Associates, Inc., Erie, Pa.; E. Severis, E. Carroll; Biostatisticians — W. Hall (head), V. Dragalin, P. Huang, K. Ding, C. Feng, J. Whitehead (consultant); Executive Committee — A. Moss (chair), M. Brown, D. Cannom, J. Daubert, W. Hall, S. Higgins, H. Klein, D. Wilber, W. Zareba.

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